

# ON THE NUMBER OF PREDECESSORS IN CONSTRAINED RANDOM MAPPINGS

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ABSTRACT. We consider random mappings from an  $n$ -element set into itself with constraints on coalescence as introduced by Arney and Bender. A local limit theorem for the distribution of the number of predecessors of a random point in such a mapping is presented by using a generating function approach and singularity analysis.

## 1. INTRODUCTION

For  $M_n = \{1, 2, \dots, n\}$  let  $F_n$  denote the set of all functions from  $M_n$  into itself equipped with the uniform distribution. Then an element of  $F_n$  is called a random mapping. Each random mapping  $f$  can be represented by a functional graph, i.e. the graph consisting of the nodes  $1, 2, \dots, n$ , and of the edges  $(i, f(i))$ ,  $i = 1, \dots, n$ . Various characteristics of random mappings have been studied. See e.g. [1, 3, 4, 6, 8, 9, 10, 12].

Arney and Bender [1] examined a more general model: They considered mappings such that the number of preimages of every point lies in a given set  $D$  of nonnegative integers (with  $0 \in D$ ) or, equivalently, the degrees of the nodes of the functional graph have to be in  $D$ . Let  $F_n^D$  denote the set of those mappings on  $M_n$ . Arney and Bender derived the distributions of many random mapping characteristics for  $F_n^D$  mainly by means of combinatorial counting arguments. In this paper we will study the distribution of a further random mapping characteristic using a generating function approach.

Let  $x \in M_n$  and  $f \in F_n^D$ .  $y \in M_n$  is called a *predecessor* of  $x$  if there exists a  $j > 0$  such that the  $j$ -th iterate of  $f$  applied on  $y$  yields  $x$ . Denote the number of predecessors of  $x$  by  $\omega(x)$ . Then we will show

**Theorem 1.** *Let  $\phi(x) = \sum_{k \in D} x^k/k!$  and  $\beta$  be the positive root of  $\beta\phi'(\beta) = \phi(\beta)$ . Furthermore define  $\lambda := \beta^2\phi''(\beta)/\phi(\beta)$  and  $d := \gcd\{k : k \in D\}$ . Then for a randomly chosen point of a mapping of  $F_n^D$  the expected number of predecessors equals  $\sim \sqrt{\frac{\pi n}{2\lambda}}$  and moreover a local limit theorem holds: If  $d|r$  then*

$$\mathbf{P}[\omega = r] = \frac{d}{\sqrt{2\pi\lambda r^3 \left(1 - \frac{r}{n}\right)}} (1 + o(1))$$

for  $r \rightarrow \infty$  and  $n - r \rightarrow \infty$ . Otherwise  $\mathbf{P}[\omega = r] = 0$

*Remark .* For the special cases  $D = \{0, 1, 2, \dots\}$  and  $D = \{0, k\}$  the above result was obtained by Rubin and Sitgreaves [11].

## 2. COMBINATORIAL BACKGROUND

The basic concept which our method relies on is that of combinatorial constructions described e.g. in [13]: Let  $\mathcal{A}$  be a set of combinatorial objects where each object  $\sigma \in \mathcal{A}$  has a size  $|\sigma|$ . If  $a_n$  denotes the number of objects in  $\mathcal{A}$  having size  $n$ , then

$$\hat{A}(z) = \sum_{n \geq 0} a_n \frac{z^n}{n!}$$

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is called the (exponential) generating function (GF) of  $\mathcal{A}$ .

The advantage of the generating function approach is the fact that there is a direct correspondence between operations on combinatorial constructions and on GFs (see [13]). Using the graph representation of random mappings it is easy to see that they may be regarded as multisets of cycles of Cayley trees, i.e. labelled rooted trees. Let  $\mathcal{A}$  denote the set of Cayley trees. Then  $\mathcal{A}$  can be defined recursively by the symbolic equation

$$\mathcal{A} = \circ * \text{mset}(\mathcal{A}),$$

where  $\circ$  represents a node. Analogously, we get for the set of random mappings,  $\mathcal{F}$ :

$$\mathcal{F} = \text{mset}(\text{cycle}(\mathcal{A}))$$

This implies the following equations for the corresponding GFs: The GF of Cayley trees is given by

$$a(z) = ze^{a(z)}$$

and the GF of random mappings

$$f(z) = \frac{1}{1-a(z)}.$$

As we want to study the random mappings with constraints on coalescence, we have to modify  $a(z)$  properly and obtain

$$a(z) = z\phi(a(z)) \quad \text{where } \phi(z) = \sum_{k \in D} \frac{z^k}{k!}.$$

and the GF of constrained random mappings is given by

$$f(z) = \frac{1}{1-b(z)} \quad \text{where } b(z) = z\phi'(\bar{a}(z))$$

since the root of each tree has an additional predecessor coming from the cycle.

Let  $c_{nk}^{(r)}$  denote the number of mappings in  $F_n^D$  which have exactly  $k$  points  $x$  satisfying  $\omega(x) = r$ , where  $\omega(x)$  is the number of predecessors of the point  $x$ . The number of all points in all mappings of  $F_n^D$  that satisfy this equation is denoted by  $b_n^{(r)}$ . Hence we have

$$\frac{1}{|F_n^D|n} b_n^{(r)} = \mathbf{P}\{\omega = r\}.$$

Thus for establishing the local limit theorem we have to compute  $b_n^{(r)}$ . Obviously the relation

$$b_n^{(r)} = \sum_{k \geq 1} k c_{nk}^{(r)}$$

holds. Therefore  $b_n^{(r)}$  can be calculated by

$$b_n^{(r)} = \left[ \frac{z^n}{n!} \right] (c_r)_u(z, 1) \tag{1}$$

where  $[z^n]f$  denotes the coefficient of  $z^n$  in the power series of  $f$  and  $(c_r)_u$  denotes the partial derivative with respect to  $u$ .  $c_r(z, u)$  is defined by

$$c_r(z, u) = \sum_{n, k \geq 0} c_{nk}^{(r)} \frac{z^n}{n!} u^k$$

and can be obtained by marking the nodes we are keeping track of in the functional graphs which corresponds to introducing a further variable in the GF. Further examples of these marking techniques in combinatorial constructions can be found in [3, 4].

(1) shows that in order to get the asymptotic behaviour of  $b_n^{(r)}$  we have to evaluate the coefficients of certain power series. In order to do this we will use the following theorem by Flajolet and Odlyzko [7].

**Theorem 2.** *Let  $f(z)$  be analytic in the domain*

$$\Delta = \{z/|z| \leq s + \eta, |\arg(z - s)| \geq \phi\},$$

where  $s, \eta$  are positive real numbers and  $0 < \phi < \frac{\pi}{2}$ . Furthermore let  $\sigma(u) = u^\alpha \log^\beta u$ ,  $\alpha \notin \{0, -1, -2, \dots\}$ . Then the Taylor coefficients of  $f$  satisfy

$$f(z) \sim \sigma\left(\frac{1}{1 - z/s}\right) \text{ for } z \rightarrow s \text{ in } \Delta \implies [z^n]f(z) \sim \frac{\sigma(n)}{s^n n \Gamma(\alpha)}.$$

Analogous formulas hold for  $\mathcal{O}$  and  $o$  instead of  $\sim$ .

### 3. PROOF OF THEOREM 1

For convenience, define  $a_r = [z^r]a(z)$ ,  $b_{mr} = [z^r]b(z)^m$  and assume that  $d = 1$ .  $b_n = [z^n]1/(1 - b(z))$ . First of all, let us set up the GF  $c_r(z, u)$ . Therefore we have to mark all nodes that are roots of a tree with size  $r$ . This leads to the following modification of  $a(z)$ :

$$t_r(z, u) = z\phi(t_r(z, u)) + (u - 1)a_r z^r.$$

Moreover we have to take into account that the set of predecessors of a node which belongs to a cycle consists of all nodes in the component. Hence in a component of size  $r$  all nodes in the cycle have to be marked. For components containing a cycle of length  $m$  this yields the GF

$$\frac{(z\phi'(t_r(z, u)))^m}{m} + \frac{1}{m}(u^m - 1)z^r b_{mr}.$$

Thus we have

$$c_r(z, u) = \frac{1}{1 - z\phi'(t_r(z, u))} \exp\left(z^r \sum_{m=1}^r \frac{u^m - 1}{m} b_{mr}\right)$$

and consequently

$$(c_r)_u(z, 1) = \frac{a_r z^{r+1} \phi''(a(z))}{(1 - b(z))^3} + \frac{z^r \sum_{m=1}^r b_{mr}}{1 - b(z)}. \quad (2)$$

Note that  $b_{mr} = 0$  if  $m = 0$  or  $m > r$  and thus

$$\sum_{m=1}^r b_{mr} = \sum_{m=0}^{\infty} b_{mr} = [z^r] \frac{1}{1 - b(z)} = b_r.$$

Hence the functions  $a(z)$  and  $b(z)$  contain the information we need.  $a(z)$  has a singularity  $\rho$  on the circle of convergence which is determined by

$$\begin{aligned} a(\rho) &= \rho\phi(a(\rho)) \\ 1 &= \rho\phi'(a(\rho)). \end{aligned}$$

To proceed we will need the following lemma which is an immediate consequence of [5, Theorem 7.1]:

**Lemma 1.** *Let  $F(z, y)$  be continuous in  $\{(x, y) : |z - z_0| < r_1, |y - y_0| < r_2\}$ . Furthermore assume that if one variable is fixed, then  $F$  is analytic in the other variable. If  $F(z_0, y_0) = F_y(z_0, y_0) = 0$ ,  $F_z(z_0, y_0) \neq 0$  and  $F_{yy}(z_0, y_0)$ , then there exists a function  $y(z)$  which admits the local representation*

$$y(z) \sim y_0 - \sqrt{\frac{2F_z(z_0, y_0)}{F_{yy}(z_0, y_0)}} \sqrt{z_0 - z}$$

for  $z \rightarrow z_0$ .

Thus in the vicinity of  $\rho$   $a(z)$  admits the following local expansion:

$$a(z) \sim \beta - \sqrt{\frac{2\phi(\beta)}{\phi''(\beta)}} \sqrt{1 - \frac{z}{\rho}} = \beta - \frac{\beta\sqrt{2}}{\sqrt{\lambda}} \sqrt{1 - \frac{z}{\rho}}, \quad \text{as } z \rightarrow \rho$$

where  $\beta = a(\rho)$  and  $\lambda = \beta^2 \phi''(\beta) / \phi(\beta)$ . By expanding  $\phi(z)$  we immediately obtain

$$b(z) \sim 1 - \rho \sqrt{2\phi(\beta)\phi''(\beta)} \sqrt{1 - \frac{z}{\rho}} = 1 - \sqrt{2\lambda} \sqrt{1 - \frac{z}{\rho}} \quad \text{as } z \rightarrow \rho$$

and thus

$$\frac{1}{1-b(z)} \sim \frac{1}{\sqrt{2\lambda}} \left( \frac{1}{1-z/\rho} \right)^{1/2} \quad \text{as } z \rightarrow \rho. \quad (3)$$

Now we are able to apply Theorem 2 and get

$$a_r \sim \frac{\beta}{\rho^r \sqrt{2\pi\lambda r^3}} \quad \text{and} \quad b_r \sim \frac{1}{\rho^r \sqrt{2\pi\lambda r}} \quad \text{as } r \rightarrow \infty. \quad (4)$$

Using this we finally get

$$\begin{aligned} \frac{b_n^{(r)}}{|F_n^D|n} &= \frac{1}{nb_n} [z^n] (c_r)_u(z, 1) \\ &\sim \frac{\rho^n \sqrt{2\pi\lambda n}}{n} \left( \frac{a_r \lambda}{\beta} [z^{n-r}] \frac{1}{(1-b(z))^3} + [z^{n-r}] \frac{1}{1-b(z)} [z^r] \frac{1}{1-b(z)} \right) \\ &\sim \frac{\rho^n \sqrt{2\pi\lambda n}}{n} \left( \frac{\lambda}{\rho^r \sqrt{2\pi\lambda r^3}} \frac{\sqrt{n-r}}{\rho^{n-r} \sqrt{2\pi\lambda^3}} + \frac{1}{\rho^r \sqrt{2\pi\lambda r}} \frac{1}{\rho^{n-r} \sqrt{2\pi\lambda(n-r)}} \right) \\ &= \frac{1}{\sqrt{2\pi\lambda r^3(1-r/n)}} \end{aligned}$$

What remains to do is the calculation of the mean value  $\mu_n$ : By (2) we have

$$\begin{aligned} \mu_n &= \frac{1}{nb_n} [z^n] \sum_{r \geq 0} r (c_r)_u(z, 1) \\ &= \frac{1}{nb_n} [z^n] \left( \frac{\phi''(a(z))}{(1-b(z))^3} z^2 a'(z) + \frac{1}{1-b(z)} z \left( \frac{1}{1-b(z)} \right)' \right) \end{aligned}$$

Using the functional equation of  $a(z)$  we immediately get

$$a'(z) = \frac{a(z)}{z(1-b(z))}.$$

This implies

$$\mu_n = \frac{1}{nb_n} [z^n] \left( \frac{2z\phi''(a(z))a(z)}{(1-b(z))^4} + \frac{z\phi'(a(z))}{(1-b(z))^3} \right)$$

Now observe that for  $z \rightarrow \rho$  we have  $z\phi''(a(z))a(z) \sim \lambda$ . Thus we get by using (3) and (4) and applying Theorem 2

$$\mu_n = \sqrt{\frac{\pi n}{2\lambda}} \left( 1 + \mathcal{O}\left(\frac{1}{\sqrt{n}}\right) \right)$$

and the proof is complete.

*Remark 1.* In an analogous way it is possible to reobtain Arney and Bender's [1] results by the method presented here or to derive local limit theorems for other random mapping characteristics. Note, for instance, that recently Baron, Drmota and Mutafchiev [2] derived the missing distribution of [1, Table II] by using a generating function approach.

*Remark 2.* It should be mentioned that the coefficient of coalescence  $\lambda$  which occurs in the distributions of several random mapping characteristics has a simple probabilistic interpretation, as Aldous pointed out: If we consider the trees which build up the random mappings as representations of Galton-Watson branching processes, then the offspring distribution is determined by the tree function  $\phi(z)$  and  $\lambda$  is equal to its variance.

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