

# REFLECTED BROWNIAN BRIDGE LOCAL TIME CONDITIONED ON ITS LOCAL TIME AT THE ORIGIN

BERNHARD GITTENBERGER\* AND GUY LOUCHARD\*\*

ABSTRACT. The moments of the local time of a reflected Brownian bridge conditioned on its local time at the origin are computed by two different methods: First, by conditioning an identity for the unconditioned local time and second, by using a limit theorem for random forests.

## 1. INTRODUCTION

**1.1. Motivating remarks.** The analysis of construction costs of hash tables under the linear probing strategy is closely related to random forests (see e.g. [15]) as well as to functionals (in particular, the occupation time) of a reflected Brownian bridge conditioned on its local time at the origin (see [9, 14, 2, 3]). We are interested in the *local version* of this functional, i.e., the local time, which appears naturally as limiting profile of random forests. This process has been described on the one hand implicitly in terms of a stochastic differential equation (see [17]), on the other hand by a weak limit theorem for the finite-dimensional distributions (see [11]). In this note, we complement these results by presenting a moment convergence theorem.

**1.2. Notations.** Throughout this note, the standard Brownian motion (BM) will be denoted by  $x(t)$ . Other classical Brownian processes are the reflected BM:  $x^+(t) := |x(t)|$ , the Brownian Bridge (BB) on  $[0, 1]$ :  $B(t)$ , the reflected BB on  $[0, 1]$ :  $B^+(t)$ . The local time of  $B^+(t)$  at level  $a$  and time  $t$  will be denoted by  $t^+(t, a)$  and we define  $t^+(t, a | b)$  to be the (total) local time at level  $a$  conditioned on having a local time at the origin equal to  $b$ .

We are interested in the moments of  $t^+(1, y | b)$ , i.e.,

$$\mu_k(y, b) := \mathbb{E} \left( t^+(1, y | b)^k \right),$$

and will present two approaches to compute these moments. One based on results of [12] and one based on a limit theorem for random forests (see [11]).

**1.3. Some known facts about normalized local time.** We will use the following abbreviations throughout the paper

$$\begin{aligned} Ch_i &:= \cosh(\sqrt{2\alpha}\rho_i) & Ch_{i,j} &:= \cosh(\sqrt{2\alpha}(\rho_i - \rho_j)) \\ Sh_i &:= \sinh(\sqrt{2\alpha}\rho_i) & Sh_{i,j} &:= \sinh(\sqrt{2\alpha}(\rho_i - \rho_j)) \\ E^i &:= e^{\sqrt{2\alpha}\rho_i} & E^{i,j} &:= e^{\sqrt{2\alpha}(\rho_i - \rho_j)} \end{aligned}$$

From [12] (we take the opportunity to correct a misprint in [12]:  $F_5$  has a  $\sqrt{2}$  in the denominator), we know that, with  $\tau^+(a) := t^+(1, a)$ ,

$$\begin{aligned} & \int_0^\infty e^{-\alpha t} \mathbb{E} \left[ e^{-\beta_1 \sqrt{t} \tau^+(\rho_1/\sqrt{t}) - \beta_2 \sqrt{t} \tau^+(\rho_2/\sqrt{t})} \right] \frac{dt}{\sqrt{2\pi t}} \\ &= \frac{\alpha}{4Ch_1^2 Sh_{2,1}^2 \left( \beta_1 + \sqrt{\frac{\alpha}{2}} \frac{Ch_2}{Ch_1 Sh_{2,1}} \right)^2 \left( \beta_2 + \sqrt{\frac{\alpha}{2}} \frac{E^{2,1}}{Sh_{2,1}} \right) - 2\alpha Ch_1^2 \left( \beta_1 + \sqrt{\frac{\alpha}{2}} \frac{Ch_2}{Ch_1 Sh_{2,1}} \right)} \end{aligned} \quad (1)$$

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\* Technische Universität Wien, Department of Geometry, Wiedner Hauptstr. 8–10/113, A-1040 Wien, Austria, email: Bernhard.Gittenberger@geometrie.tuwien.ac.at.

\*\* Université Libre de Bruxelles, Département d'Informatique, CP 212, Boulevard du Triomphe, B-1050 Bruxelles, Belgium, email:louchard@ulb.ac.be.

where  $\beta_1$  and  $\beta_2$  are arbitrary positive constants.

*Remark.* Note that since  $\rho_1 < \rho_2$ , the transform cannot be symmetric with respect to  $\beta_1$  and  $\beta_2$ .

In order to find the distribution of the BB local time at the origin we may use the results in Louchard [13] or the path transformation techniques in Bertoin and Pitman [1]. But for illustration we use  $t^+(t, x) \stackrel{\mathcal{D}}{=} \sqrt{t}t^+(1, x/\sqrt{t})$  in conjunction with Equ. (1) to get

$$\mathbb{E}_0 \int_0^\infty e^{-\alpha t} \left[ e^{-\beta t^+(t,0)} \Big| x(t) = 0 \right] \frac{1}{\sqrt{2\pi t}} dt = \frac{1}{2\beta + \sqrt{2\alpha}}.$$

Inverting on  $\beta$  leads to

$$\int_0^\infty e^{-\alpha t} \text{Pr}_0[[t^+(t,0) | x(t) = 0] \in db] \frac{1}{\sqrt{2\pi t}} dt = e^{-\sqrt{2\alpha}b/2} db$$

or

$$\text{Pr}_0[[t^+(t,0) | x(t) = 0] \in db] = \frac{be^{-b^2/(8t)} db}{(4t)}$$

which is nothing but the Rayleigh density.

**1.4. Some known facts about random forests.** We consider the set  $F(n, N)$  of random forests consisting of  $n$  vertices and  $N$  rooted trees which can be viewed as realizations of Galton-Watson branching processes with  $N$  initial particles and conditioned to have total progeny  $n$ . This means that we start with  $N$  particles each of which produces children. Then each of the children produces further descendants, and so on. The number of children of a particle is governed by a given distribution and all particles (of the whole tree) behave independently (i.e., their offsprings are *iid* random variables). Let  $b_{n,N}$  denote the number of forests in  $F(n, N)$ , weighted according to the probability on  $F(n, N)$ , which is the above described probability conditioned on the total progeny. It is well known (see e.g. [16]) that the generating function for those forests is  $b(z) = a(z)^N$  with

$$a(z) = z\phi(a(z)). \quad (2)$$

Here  $a(z)$  is the generating function for a single tree. Let  $L_{n,N}(k)$  be the number of vertices at height  $k$  in a random forest in  $F(n, N)$ .

By [11, Theorem 1.3] we have for  $n, N \rightarrow \infty$  such that  $2N/\sigma\sqrt{n} \rightarrow \alpha > 0$

$$\mathbb{E} \left( \frac{2L_{n,N}(2\kappa\sqrt{n})}{\sigma\sqrt{n}} \right)^d \rightarrow \mathbb{E} (t^+(1, \kappa|\alpha)^d) \quad (3)$$

This can be used to compute the moments  $\mu_k(y, b)$ .

## 2. THE MOMENTS

**Theorem 1.** *The moments of the local time of the reflected Brownian bridge conditioned to have local time  $b$  at the origin satisfy*

$$\mu_k(y, b) = 2\sqrt{2\pi}k!e^{b^2/8} \sum_{v=0}^{k-1} \frac{(-1)^v}{(k-v)!} \binom{k-1}{v} b^{k-v-1} \sum_{j=0}^v \binom{v}{j} (-1)^j \phi^{(v-1)} \left( -\frac{b+4(k-j)y}{2} \right) \quad (4)$$

where  $\phi^{(1)}(x) = \phi(x)$  denotes the Gaussian distribution function and  $\phi^{(j+1)}(x) := \int_{-\infty}^x \phi^{(j)}(u) du$ .  $\phi^{(0)}(x)$  denotes the density of the standard normal distribution and  $\phi^{(-1)}(x)$  its derivative.

**2.1. First approach.** In the same manner as in section 1.3 we analyze the reflected BB local time at  $x$ , conditioned on the local time at 0. Equ. (1) leads to

$$\begin{aligned} \mathbb{E}_0 \int_0^\infty dt e^{-\alpha t} \int_0^\infty db e^{-\sqrt{t}\tau^+(\rho_1/\sqrt{t}|b)} b e^{-b^2/8} e^{-\delta\sqrt{t}b} \frac{1}{4\sqrt{2\pi t}} \\ = \frac{\alpha}{4Sh_1^2 \left( \delta + \sqrt{\frac{\alpha}{2}} \frac{Ch_1}{Sh_1} \right)^2 \left( 1 + \sqrt{\frac{\alpha}{2}} \frac{E^1}{Sh_1} \right) - 2\alpha \left( \delta + \sqrt{\frac{\alpha}{2}} \frac{Ch_1}{Sh_1} \right)}. \end{aligned} \quad (5)$$

On setting  $v := \sqrt{tb}$  the left-hand side becomes

$$\int_0^\infty db \int_0^\infty dv e^{-\alpha v^2/b^2} \mathbb{E} \left[ e^{-v/b\tau^+(\rho_1 b/v|b)} \right] e^{-b^2/8} e^{-\delta v} \frac{1}{2\sqrt{2\pi}}$$

The right-hand side is of the form

$$\frac{A}{B(\delta + D)^2 + C(\delta + D)} = \left[ \frac{A}{B} \cdot \frac{1}{s} \cdot \frac{1}{s + C/B} \right]_{s=\delta+D}$$

and thus inversion on  $\delta$  using standard rules as listed, for example, in [6] yields

$$\frac{A}{C} \left( e^{-Dv} - e^{-v(C+BD)/B} \right).$$

Hence we obtain from (5)

$$\int_0^\infty db e^{-b^2/8} e^{-\alpha v^2/b^2} \mathbb{E} \left[ e^{-v/b\tau^+(\rho_1 b/v|b)} \right] \frac{1}{2\sqrt{2\pi}} = F(v, \alpha, \rho_1, b) \quad (6)$$

with

$$F(v, \alpha, \rho_1, b) = \frac{1}{2} \left[ \exp \left( -v \sqrt{\frac{\alpha}{2}} \frac{Ch_1 + \sqrt{\frac{\alpha}{2}} \frac{E^1 Ch_1 - 1}{Sh_1}} \right) - \exp \left( -v \sqrt{\frac{\alpha}{2}} \frac{Ch_1}{Sh_1} \right) \right].$$

Substitute  $w = \alpha v^2, \eta = \alpha^{-1/2}, y = \rho_1 b/v$ , and  $s := 1/b^2$ . Then equ. (6) becomes

$$\int_0^\infty e^{-1/(8s)} e^{-ws} \mathbb{E} \left[ e^{-\eta \sqrt{ws} \tau^+(y|b)} \right] \frac{ds}{4s^{3/2} \sqrt{2\pi}} = G(w, \eta) \quad (7)$$

with

$$G(w, \eta) = G(w, \eta, y, s) := F \left( \sqrt{\frac{w}{\alpha}}, \frac{1}{\eta^2}, \frac{yv}{b}, \frac{1}{\sqrt{s}} \right)$$

Expanding both sides of (7) w.r.t  $\eta$ , this gives

$$\int_0^\infty e^{-1/(8s)} e^{-ws} (-1)^k s^{k/2} \mu_k(y, b) \frac{ds}{4s^{3/2} \sqrt{2\pi} k!} = [\eta^k] \frac{G(w, \eta)}{w^{k/2}}. \quad (8)$$

The LHS of (8),  $L_1$  say, is a Laplace transform, the RHS,  $R_1(k)$ , say, can be computed as follows. To simplify the notations set

$$X = \exp \left( -\sqrt{\frac{w}{2}} \right), \quad Y = \exp \left( -\frac{2\sqrt{2wy}}{b} \right).$$

$R_1(k)$  becomes

$$[\eta^k] \frac{G_1(w, \eta)}{w^{k/2}}$$

with

$$\begin{aligned} G_1(w, \eta) &= \frac{1}{2} \exp \left( -\frac{\sqrt{w}}{\sqrt{2}} \frac{1 + \eta(1+Y)/\sqrt{2}}{1 + \eta(1-Y)/\sqrt{2}} \right) \\ &= \frac{X}{2} \exp \left( -\frac{\sqrt{w}}{\sqrt{2}} \frac{2\eta Y/\sqrt{2}}{1 + \eta(1-Y)/\sqrt{2}} \right) \\ &= \frac{X}{2} \exp \left( \frac{\sqrt{w}}{\sqrt{2}} \left[ \sum_1^\infty (-\eta/\sqrt{2})^i (1-Y)^{i-1} 2Y \right] \right) \end{aligned}$$

Setting  $z := \sqrt{w}$  and expanding this (omitting the detailed derivation)

$$R_1(k) = \frac{X}{2} (-1)^k \sum_{v=0}^{k-1} \frac{(-1)^v}{(k-v)!} \frac{1}{(\sqrt{2}z)^v} \binom{k-1}{v} \sum_{j=0}^v \binom{v}{j} (-1)^j Y^{k-j} \quad (9)$$

Clearly, we need to invert Laplace transforms of the form  $e^{-\sqrt{\gamma w}/w^{(j+1)/2}}$ . But in [4] the following lemma is proved:

**Lemma 1.** *We have*

$$\int_0^\infty \phi^{(j)}(-b) e^{-ws} \frac{(2s)^{(j+1)/2}}{s} ds = \frac{e^{-\sqrt{2w}}}{w^{(j+1)/2}}, \quad j \geq 1$$

where  $b = 1/\sqrt{s}$ .

It is also well known that

$$\begin{aligned} \int_0^\infty e^{-ws} e^{-1/(2s)} \frac{ds}{\sqrt{2\pi s^{3/2}}} &= e^{-\sqrt{2w}}, \\ \int_0^\infty e^{-ws} e^{-1/(2s)} \frac{ds}{\sqrt{\pi s}} &= e^{-\sqrt{2w}}/\sqrt{w}. \end{aligned}$$

Let us remark that  $\phi^{(j)}(x)$  can be expressed in term of  $\phi(x)$  and  $e^{-x^2/2}/\sqrt{2\pi}$  in the form

$$\phi^{(j)}(x) = p_{1,j}(x)\phi(x) + p_{2,j}(x)e^{-x^2/2}/\sqrt{2\pi},$$

where  $p_{1,j}(x)$  and  $p_{2,j}(x)$  are polynomials of degree  $j-1$  and  $j-2$ , respectively. In [4], we provide an efficient procedure to compute  $\phi^{(j)}(x)$ . Standard transformations lead to formulae inverting expressions of the form  $e^{-\sqrt{\gamma w}}/w^{(j+1)/2}$ . Now we note that each term of type  $XY^l$  in  $R_1(\cdot)$  leads to

$$\sqrt{\gamma} = \sqrt{2}/2 + l2\sqrt{2}y/b$$

Inverting (8) is now trivial (with Maple!). This gives, for instance,

$$\mu_1(y, b) = (b + 4y)e^{-y(b+2y)} \quad (10)$$

$$\mu_2(y, b) = \left(4e^{-(4y+b)^2/8} + (b^2 + 8by - 4)e^{-(8y+b)^2/8}\right) e^{b^2/8} \quad (11)$$

$$\begin{aligned} \mu_3(y, b) &= \left(12\sqrt{2\pi} \left(\phi\left(-\frac{b+4y}{2}\right) - 2\phi\left(-\frac{b+8y}{2}\right) + \phi\left(-\frac{b+12y}{2}\right)\right)\right) \\ &\quad + 12be^{-(8y+b)^2/8} + (b^3 + 12b^2y - 12b)e^{-(12y+b)^2/8} \Big) e^{b^2/8} \quad (12) \end{aligned}$$

and, finally, application on (9) leads to (4).

*Remark 1.* As a check, we may compute  $\mathbb{E} \int_0^1 B^+(t|b) dt = \int_0^\infty \mathbb{E}[\tau^+(x|b)] x dx$  and get exactly the expression we obtained in [4],  $\phi(-b/2)\sqrt{2\pi}e^{b^2/8}/2$  (the local time at the origin used here is twice the local time at the origin considered in [4]).

*Remark 2.* Integrating the moments w.r.t. the  $b$  density gives the unconditioned moments as computed by Takacs in [18]. E.g., Maple gives

$$\begin{aligned} \mu_1(y) &= 2\sqrt{2\pi}(1 - \phi(2y)) \\ \mu_2(y) &= -24y\sqrt{2\pi} + 4e^{-2y^2} + 4e^{-8y^2} + 16y\sqrt{2\pi}\phi(4y) + 8y\sqrt{2\pi}\phi(2y) \end{aligned}$$

higher moments can be integrated with Maple only numerically.

**2.2. Second approach.** Let  $b_{k,m,n,N}$  denote the (weighted) number of forests in  $F(n, N)$  with  $m$  nodes in stratum  $k$ . Then standard methods on generating functions (see [10] for a general introduction and [8] for the treatment of the particular case of random trees) give

$$\mathbf{P}\{L_{n,N}(k) = m\} = \frac{b_{k,m,n,N}}{b_{n,N}} = \frac{1}{b_{n,N}} [z^n u^m] y_k(z, ua(z))^N$$

where

$$y_0(z, u) = u, \quad y_{i+1}(z, u) = z\varphi(y_i(z, u)), \quad i \geq 0.$$

Consequently, we have

$$\mathbb{E}L_{n,N}(k) = \frac{1}{b_{n,N}} [z^n] \frac{\partial}{\partial u} y_k(z, ua(z))^N \Big|_{u=1}$$

and moreover

$$\left(\frac{2}{\sigma\sqrt{n}}\right)^m \mathbb{E}L_{n,N}(k)(L_{n,N}(k) - 1) \cdots (L_{n,N}(k) - m + 1) \sim \left(\frac{2}{\sigma\sqrt{n}}\right)^m \mathbb{E}L_{n,N}(k)^m.$$

Thus by (3) the moments can be calculated by

$$\mu_k(y, b) = \lim_{n \rightarrow \infty} \left( \frac{2}{\sigma\sqrt{n}} \right)^k \frac{1}{b_{n,N}} [z^n] \frac{\partial^k}{\partial u^k} y_\ell(z, ua(z))^N \Big|_{u=1} \quad (13)$$

with  $2N/\sigma\sqrt{n} \sim b$  and  $\ell = \lfloor 2y\sqrt{n}/\sigma \rfloor$ . The calculation of these coefficients is done by singularity analysis and contour integration. The relevant asymptotic expansions are provided in the following lemma:

**Lemma 2.** *Let  $a = a(z)$  be the tree function (2) and  $\alpha = \alpha(z) := z\varphi'(a(z))$ . Furthermore,  $z_0$  denotes the dominant singularity of  $a(z)$  and  $S := a(z_0)$ . Then for  $b = 2N/\sigma\sqrt{n}$  and  $\ell = 2y\sqrt{n}/\sigma$  the following asymptotic expansion holds as  $z \rightarrow z_0$  such that  $z - z_0 \notin \mathbb{R}^+$ :*

$$[z^n] \frac{1}{(1-\alpha)^m} \alpha^\ell a^N \sim \frac{S^N n^{(m-2)/2}}{\sigma^m z_0^n} \phi^{(m-1)} \left( -\frac{b+4y}{2} \right)$$

where  $m$  is a fixed nonnegative integer.

*Proof.* The lemma can be proved by an appropriate modification of the proof of [7, Theorem 4]  $\square$

When computing the derivatives in (13) only the main term of  $\frac{\partial^k}{\partial u^k} y_\ell(z, ua(z))$  is relevant for the calculation of  $\mu_k(y, b)$ . Set  $\beta = \beta(z) := z\varphi''(a(z))$ . Then, using  $\alpha(z) \sim 1$  for  $z \rightarrow z_0$  (see [16] for a detailed expansion), we get by elementary calculations

$$\begin{aligned} \frac{\partial}{\partial u} y_\ell(z, u) \Big|_{u=a(z)} &= \alpha^\ell \\ \frac{\partial^2}{\partial u^2} y_\ell(z, ua(z)) \Big|_{u=a(z)} &= \beta \alpha^{\ell-1} \frac{1-\alpha^\ell}{1-\alpha} \\ \frac{\partial^3}{\partial u^3} y_\ell(z, u) \Big|_{u=a(z)} &\sim 3\beta^2 \alpha^{\ell-1} \frac{(1-\alpha^\ell)(1-\alpha^{\ell-1})}{(1-\alpha)(1-\alpha^2)} \end{aligned}$$

Plugging these formulas into  $\frac{\partial^j}{\partial u^j} y_\ell(z, ua(z))^N$ , using  $b_{n,N} \sim NS^N e^{-N^2/2m\sigma^2} / \sigma z_0^n \sqrt{2\pi n^3}$  and applying Lemma 2 gives (10)–(12).

This leads to the following lemma

**Lemma 3.** *If  $z \rightarrow z_0$  such that  $z - z_0 \notin \mathbb{R}^+$ , then the following expansion holds:*

$$\frac{\partial^m}{\partial u^m} y_\ell(z, u) \Big|_{u=a(z)} \sim \frac{m! \beta^{m-1} \alpha^\ell}{2^{m-1}} \left( \frac{1-\alpha^\ell}{1-\alpha} \right)^{m-1}$$

*Proof.* Note that Faà di Bruno's formula (see e.g. [5]) gives

$$\frac{\partial^{m+1} y_\ell}{\partial u^{m+1}}(z, 1) = \sum_{\sum_{i=1}^m i k_i = m+1} \frac{(m+1)!}{k_1! \cdots k_m!} z\varphi^{(k_1+\cdots+k_m)}(a(z)) \prod_{j=1}^m \left( \frac{1}{j!} \frac{\partial^j y_{\ell-1}}{\partial u^j} \right)^{k_j} + \alpha(z) \frac{\partial^{m+1} y_{\ell-1}}{\partial u^{m+1}}(z, 1).$$

By induction and using  $\sum (j-1)k_j = m+1 - \sum k_j$  we get

$$\prod_{j=1}^m \left( \frac{1}{j!} \frac{\partial^j y_{\ell-1}}{\partial u^j} \right)^{k_j} \sim \left( \frac{\beta}{2} \right)^{m+1-\sum k_j} (\alpha^{\ell-1})^{\sum k_j} \left( \frac{1-\alpha^{\ell-1}}{1-\alpha} \right)^{m+1-\sum k_j}.$$

The dominant ones of these terms are clearly those where  $m+1 - \sum k_j$  is maximal, which occurs if and only if  $\sum k_j = 2$ . This is equivalent to the existence of an  $i$  such that  $k_\lambda = 1$  if  $\lambda \in \{i, m+1-i\}$  and 0 else. In these cases we have  $z\varphi^{(k_1+\cdots+k_m)}(a(z)) = \beta$  and thus

$$\begin{aligned} \frac{\partial^{m+1} y_\ell}{\partial u^{m+1}}(z, 1) &\sim \frac{(m+1)! \beta^m m}{2^m} \alpha^{2\ell-2} \left( \frac{1-\alpha^{\ell-1}}{1-\alpha} \right)^{m-1} + \alpha(z) \frac{\partial^{m+1} y_{\ell-1}}{\partial u^{m+1}}(z, 1) \\ &\sim \frac{(m+1)! \beta^m \alpha^\ell}{2^m (1-\alpha)^{m-1}} m \sum_{i=1}^{\ell-1} \alpha^{\ell-i-1} (1-\alpha^{\ell-i})^{m-1}. \end{aligned}$$

Using the binomial theorem and summing up w.r.t  $\ell$  gives

$$\begin{aligned} \frac{\partial^{m+1} y_\ell}{\partial u^{m+1}}(z, 1) &\sim \frac{(m+1)! \beta^m \alpha^\ell}{2^m (1-\alpha)^m} m \sum_{j=0}^{m-1} \binom{m-1}{j} (-1)^j \alpha^j \frac{1 - (\alpha^{\ell-1})^{j+1}}{j+1} \\ &= \frac{(m+1)! \beta^m \alpha^\ell}{2^m (1-\alpha)^m} \left( (1 - \alpha^{\ell-1})^m - \frac{(1-\alpha)^m}{\alpha} \right) \sim \frac{(m+1)! \beta^m \alpha^\ell}{2^m} \left( \frac{1 - \alpha^\ell}{1 - \alpha} \right)^m \end{aligned}$$

where we used  $\alpha \sim 1$  in the last step.  $\square$

With the help of this lemma we can prove (4) now. By Faà di Bruno's formula we get

$$\left( \frac{\partial}{\partial u} \right)^m y_\ell^N = \sum_{\sum ik_i=m} \frac{m!}{k_1! \cdots k_m!} N(N-1) \cdots (N - \sum k_i + 1) y_\ell^{N - \sum k_i} \prod_{j=1}^m \left( \frac{1}{j!} \frac{\partial^j y_\ell}{\partial u^j} \right)^{k_j}.$$

Inserting  $u = a(z)$  and using Lemma 3 we obtain

$$\begin{aligned} \left( \frac{\partial}{\partial u} \right)^m y_\ell^N &\sim \sum_{\sum ik_i=m} \frac{m!}{k_1! \cdots k_m!} N(N-1) \cdots (N - \sum k_i + 1) \alpha^{N + \sum (i-1)k_i} \prod_{j=1}^m \left( \frac{\beta^{j-1} \alpha^\ell}{2^{j-1}} \left( \frac{1 - \alpha^\ell}{1 - \alpha} \right)^{j-1} \right)^{k_j} \\ &\sim \sum_{\sum ik_i=m} \frac{m!}{k_1! \cdots k_m!} N^{\sum k_i} \left( \frac{\beta}{2} \right)^{m - \sum k_i} (\alpha^\ell)^{\sum k_i} \left( \frac{1 - \alpha^\ell}{1 - \alpha} \right)^{m - \sum k_i} \alpha^{N + m - \sum k_i} \end{aligned}$$

Now apply Lemma 2 and use  $\beta \sim \sigma^2/S$  and set as before  $2N/\sigma\sqrt{n} = b$  and  $\ell = \lfloor 2y\sqrt{n}/\sigma \rfloor$ . This gives

$$[z^n] \left( \frac{\partial}{\partial u} \right)^m y_\ell^N \sim \sum_{\sum ik_i=m} \frac{m!}{k_1! \cdots k_m!} f\left(\sum k_i\right) \quad (14)$$

where

$$f(k) = \frac{n^{(m-2)/2} S^N}{z_0^n 2^m} b^k \sum_{i=0}^{m-k} \binom{m-k}{i} (-1)^i \phi^{(m-k-1)} \left( -\frac{b + 4(k+i)y}{2} \right).$$

Rewriting (14) in the form

$$[z^n] \left( \frac{\partial}{\partial u} \right)^m y_\ell^N \sim \sum_{k=1}^{m-1} f(k) \sum_{\sum ik_i=m, \sum k_i=k} \frac{m!}{k_1! \cdots k_m!},$$

applying (see [5, p. 135])

$$\sum_{\sum ik_i=m, \sum k_i=k} \frac{m!}{k_1! \cdots k_m!} = \binom{m-1}{m-k} \frac{m!}{k!},$$

and substituting  $k = m - v$  yields finally (4) and completes the proof.

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